

Definitions of ψ -Functions Available in Robustbase

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Preamble

Unless otherwise stated, the following definitions of functions are given by [Maronna et al. \(2006, p. 31\)](#), however our definitions differ sometimes slightly from theirs, as we prefer a different way of *standardizing* the functions. To avoid confusion, we first define ψ - and ρ -functions.

Definition 1 A ψ -function is a piecewise continuous function $\psi : \mathbb{R} \rightarrow \mathbb{R}$ such that

1. ψ is odd, i.e., $\psi(-x) = -\psi(x) \forall x$,
2. $\psi(x) \geq 0$ for $x \geq 0$, and $\psi(x) > 0$ for $0 < x < x_r := \sup\{\tilde{x} : \psi(\tilde{x}) > 0\}$ ($x_r > 0$, possibly $x_r = \infty$).
- 3* Its slope is 1 at 0, i.e., $\psi'(0) = 1$.

Note that ‘3*’ is not strictly required mathematically, but we use it for standardization in those cases where ψ is continuous at 0. Then, it also follows (from 1.) that $\psi(0) = 0$, and we require $\psi(0) = 0$ also for the case where ψ is discontinuous in 0, as it is, e.g., for the M-estimator defining the median.

Definition 2 A ρ -function can be represented by the following integral of a ψ -function,

$$\rho(x) = \int_0^x \psi(u) du, \quad (1)$$

which entails that $\rho(0) = 0$ and ρ is an even function.

A ψ -function is called *redescending* if $\psi(x) = 0$ for all $x \geq x_r$ for $x_r < \infty$, and x_r is often called *rejection point*. Corresponding to a redescending ψ -function, we define the function $\tilde{\rho}$, a version of ρ standardized such as to attain maximum value one. Formally,

$$\tilde{\rho}(x) = \rho(x)/\rho(\infty). \quad (2)$$

Note that $\rho(\infty) = \rho(x_r) \equiv \rho(x) \forall |x| > x_r$. $\tilde{\rho}$ is a ρ -function as defined in [Maronna et al. \(2006\)](#) and has been called χ function in other contexts. For example, in package `robustbase`, `Mchi(x, *)` computes $\tilde{\rho}(x)$, whereas `Mpsi(x, *, deriv=-1)` (“(-1)-st derivative” is the primitive or antiderivative) computes $\rho(x)$, both according to the above definitions.

Note: An alternative slightly more general definition of *redescending* would only require $\rho(\infty) := \lim_{x \rightarrow \infty} \rho(x)$ to be finite. E.g., “`Welsh`” does *not* have a finite rejection point, but *does* have bounded ρ , and hence well defined $\rho(\infty)$, and we *can* use it in `lmrob()`.¹

Weakly redescending ψ functions. Note that the above definition does require a finite rejection point x_r . Consequently, e.g., the score function $s(x) = -f'(x)/f(x)$ for the Cauchy ($= t_1$) distribution, which is $s(x) = 2x/(1+x^2)$ and hence non-monotone and “re descends” to 0 for $x \rightarrow \pm\infty$, and $\psi_C(x) := s(x)/2$ also fulfills $\psi_C'(0) = 1$, but it has $x_r = \infty$ and hence $\psi_C()$ is *not* a redescending ψ -function in our sense. As they appear e.g. in the MLE for t_ν , we call ψ -functions fulfilling $\lim_{x \rightarrow \infty} \psi(x) = 0$ *weakly redescending*. Note that they’d naturally fall into two sub categories, namely the one with a *finite* ρ -limit, i.e. $\rho(\infty) := \lim_{x \rightarrow \infty} \rho(x)$, and those, as e.g., the t_ν score functions above, for which $\rho(x)$ is unbounded even though $\rho' = \psi$ tends to zero.

1 Monotone ψ -Functions

Monotone ψ -functions lead to convex ρ -functions such that the corresponding M-estimators are defined uniquely.

Historically, the “Huber function” has been the first ψ -function, proposed by Peter Huber in [Huber \(1964\)](#).

¹E-mail Oct. 18, 2014 to Manuel and Werner, proposing to change the definition of “redescending”.

1.1 Huber

The family of Huber functions is defined as,

$$\rho_k(x) = \begin{cases} \frac{1}{2}x^2 & \text{if } |x| \leq k \\ k(|x| - \frac{k}{2}) & \text{if } |x| > k \end{cases},$$

$$\psi_k(x) = \begin{cases} x & \text{if } |x| \leq k \\ k \operatorname{sign}(x) & \text{if } |x| > k \end{cases}.$$

The constant k for 95% efficiency of the regression estimator is 1.345.

```
> plot(huberPsi, x., ylim=c(-1.4, 5), leg.loc="topright", main=FALSE)
```

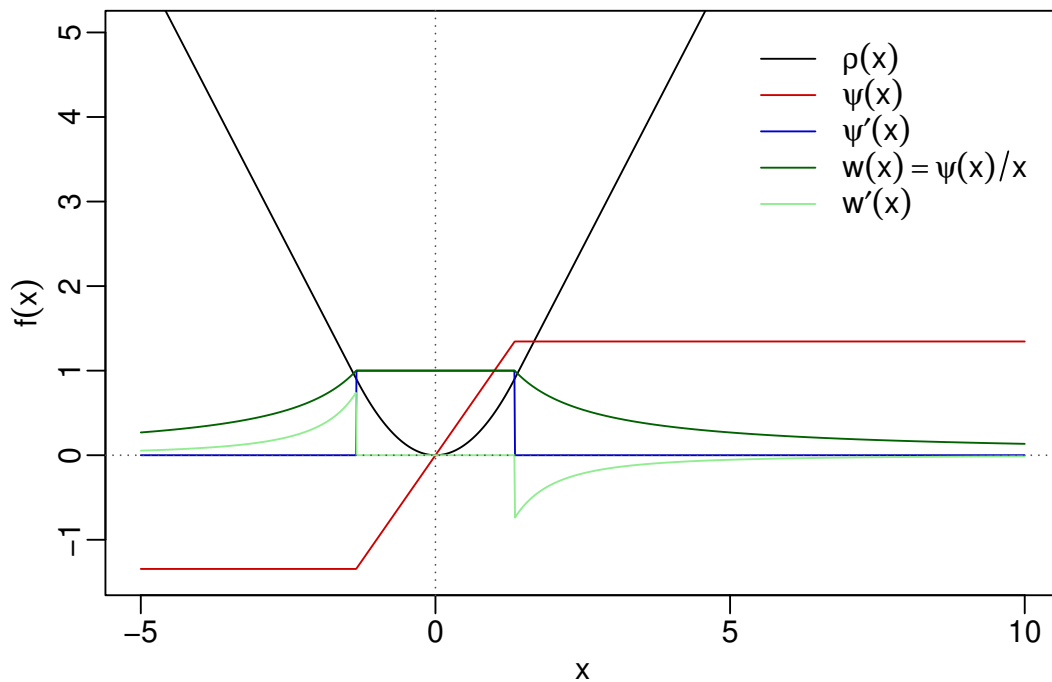


Figure 1: Huber family of functions using tuning parameter $k = 1.345$.

2 Redescenders

For the MM-estimators and their generalizations available via `lmrob()` (and for some methods of `nlrob()`), the ψ -functions are all redescending, i.e., with finite “rejection point” $x_r = \sup\{t; \psi(t) > 0\} < \infty$. From `lmrob`, the psi functions are available via `lmrob.control`, or more directly, `.Mpsi.tuning.defaults`,

```
> names(.Mpsi.tuning.defaults)
```

```
[1] "huber"      "bisquare"  "welsh"     "ggw"       "lqq"
[6] "optimal"   "hampel"
```

and their ψ , ρ , ψ' , and weight function $w(x) := \psi(x)/x$, are all computed efficiently via C code, and are defined and visualized in the following subsections.

2.1 Bisquare

Tukey's bisquare (aka "biweight") family of functions is defined as,

$$\tilde{\rho}_k(x) = \begin{cases} 1 - \left(1 - \left(\frac{x}{k}\right)^2\right)^3 & \text{if } |x| \leq k \\ 1 & \text{if } |x| > k \end{cases},$$

with derivative $\tilde{\rho}'_k(x) = 6\psi_k(x)/k^2$ where,

$$\psi_k(x) = x \left(1 - \left(\frac{x}{k}\right)^2\right)^2 \cdot I_{\{|x| \leq k\}}.$$

The constant k for 95% efficiency of the regression estimator is 4.685 and the constant for a breakdown point of 0.5 of the S-estimator is 1.548. Note that the *exact* default tuning constants for M- and MM- estimation in `robustbase` are available via `.Mpsi.tuning.default()` and `.Mchi.tuning.default()`, respectively, e.g., here,

```
> print(c(k.M = .Mpsi.tuning.default("bisquare"),  
+        k.S = .Mchi.tuning.default("bisquare")), digits = 10)
```

```
      k.M      k.S  
4.685061 1.547640
```

and that the `p.psiFun(.)` utility is available via

```
> source(system.file("xtraR/plot-psiFun.R", package = "robustbase", mustWork=TRUE))
```

```
> p.psiFun(x., "biweight", par = 4.685)
```

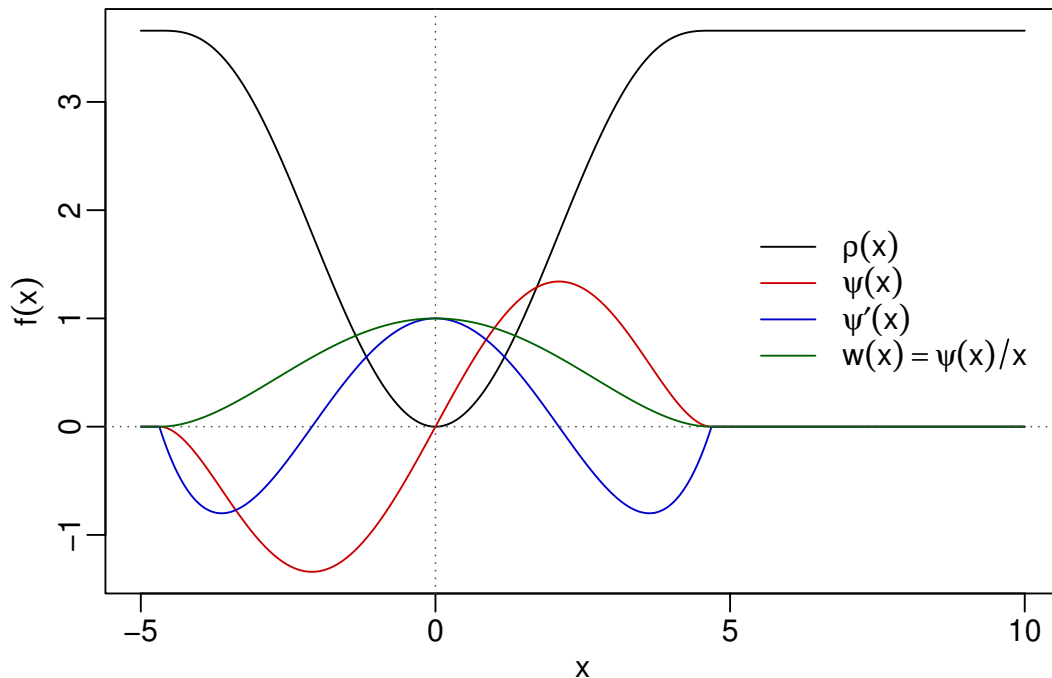


Figure 2: Bisquare family functions using tuning parameter $k = 4.685$.

2.2 Hampel

The Hampel family of functions (Hampel et al., 1986) is defined as,

$$\tilde{\rho}_{a,b,r}(x) = \begin{cases} \frac{1}{2}x^2/C & |x| \leq a \\ \left(\frac{1}{2}a^2 + a(|x| - a)\right)/C & a < |x| \leq b \\ \frac{a}{2} \left(2b - a + (|x| - b) \left(1 + \frac{r-|x|}{r-b}\right)\right)/C & b < |x| \leq r \\ 1 & r < |x| \end{cases},$$

$$\psi_{a,b,r}(x) = \begin{cases} x & |x| \leq a \\ a \operatorname{sign}(x) & a < |x| \leq b \\ a \operatorname{sign}(x) \frac{r-|x|}{r-b} & b < |x| \leq r \\ 0 & r < |x| \end{cases},$$

where $C := \rho(\infty) = \rho(r) = \frac{a}{2}(2b - a + (r - b)) = \frac{a}{2}(b - a + r)$.

As per our standardization, ψ has slope 1 in the center. The slope of the redescending part ($x \in [b, r]$) is $-a/(r - b)$. If it is set to $-\frac{1}{2}$, as recommended sometimes, one has

$$r = 2a + b.$$

Here however, we restrict ourselves to $a = 1.5k$, $b = 3.5k$, and $r = 8k$, hence a redescending slope of $-\frac{1}{3}$, and vary k to get the desired efficiency or breakdown point.

The constant k for 95% efficiency of the regression estimator is 0.902 (0.9016085, to be exact) and the one for a breakdown point of 0.5 of the S-estimator is 0.212 (i.e., 0.2119163).

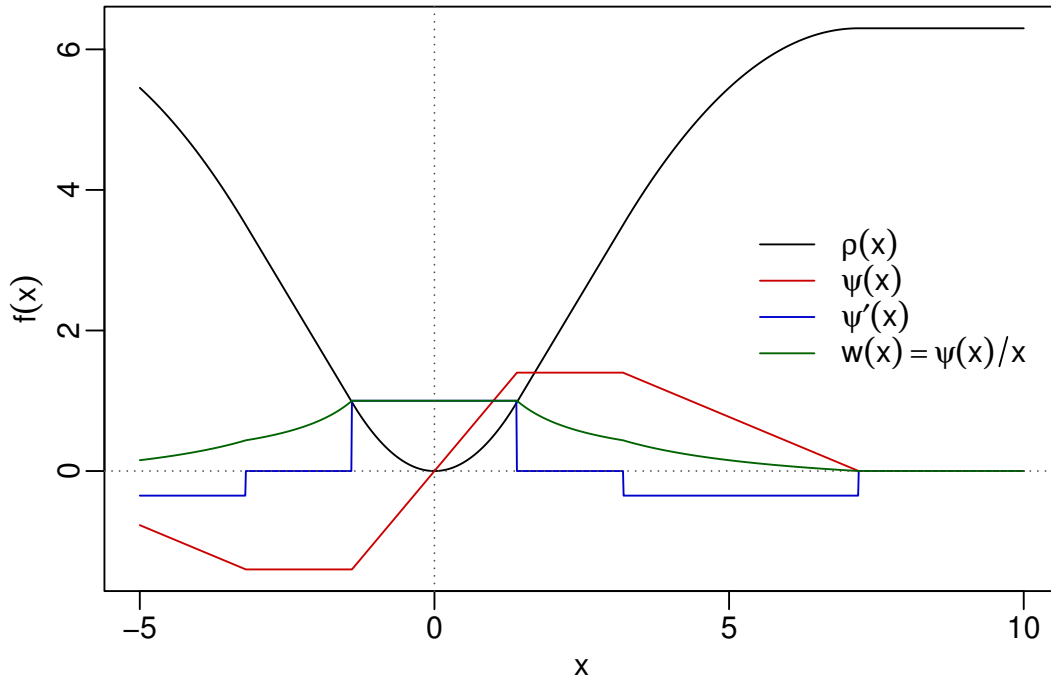


Figure 3: Hampel family of functions using tuning parameters $0.902 \cdot (1.5, 3.5, 8)$.

2.3 GGW

The Generalized Gauss-Weight function, or *ggw* for short, is a generalization of the Welsh ψ -function (subsection 2.6). In [Koller and Stahel \(2011\)](#) it is defined as,

$$\psi_{a,b,c}(x) = \begin{cases} x & |x| \leq c \\ \exp\left(-\frac{1}{2}\left(\frac{|x|-c}{a}\right)^b\right)x & |x| > c \end{cases} .$$

Our constants, fixing $b = 1.5$, and minimal slope at $-\frac{1}{2}$, for 95% efficiency of the regression estimator are $a = 1.387$, $b = 1.5$ and $c = 1.063$, and those for a breakdown point of 0.5 of the S-estimator are $a = 0.204$, $b = 1.5$ and $c = 0.296$:

```
> cT <- rbind(cc1 = .psi.ggw.findc(ms = -0.5, b = 1.5, eff = 0.95      ),
+            cc2 = .psi.ggw.findc(ms = -0.5, b = 1.5,                bp = 0.50)); cT

      [,1]      [,2] [,3]      [,4]      [,5]
cc1    0 1.3863620  1.5 1.0628199 4.7773893
cc2    0 0.2036739  1.5 0.2959131 0.3703396
```

Note that above, $cc*[1] = 0$, $cc*[5] = \rho(\infty)$, and $cc*[2:4] = (a, b, c)$. To get this from (a, b, c) , you could use

```
> ipsi.ggw <- .psi2ipsi("GGW") # = 5
> ccc <- c(0, cT[1, 2:4], 1)
> integrate(.Mpsi, 0, Inf, ccc=ccc, ipsi=ipsi.ggw)$value # = rho(Inf)
```

```
[1] 4.777389
```

```
> p.psiFun(x., "GGW", par = c(-.5, 1, .95, NA))
```

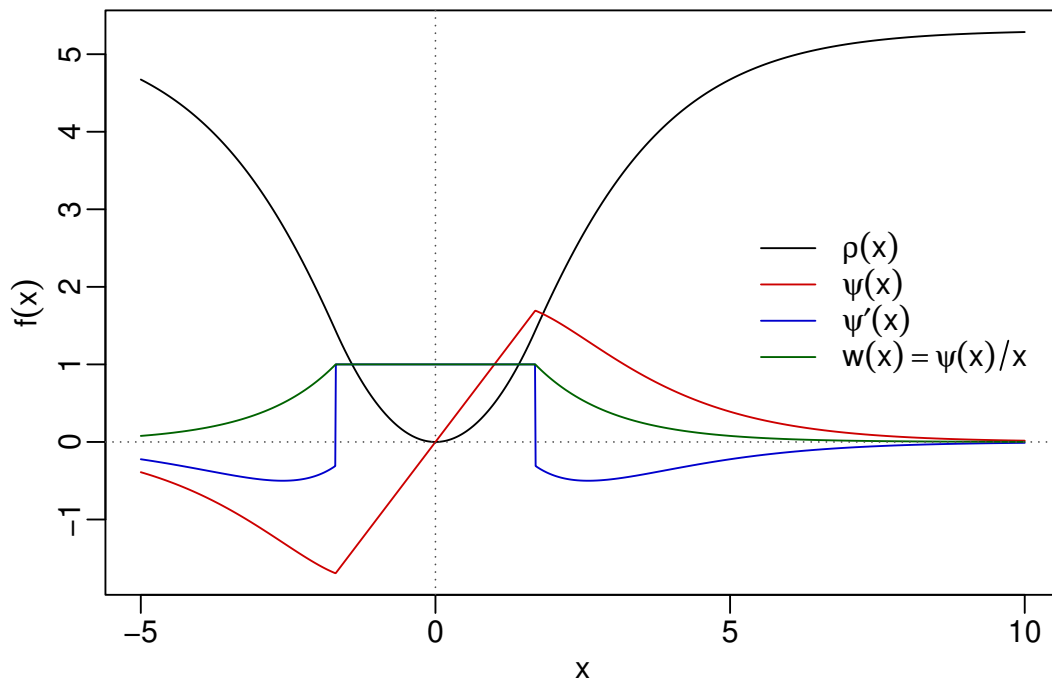


Figure 4: GGW family of functions using tuning parameters $a = 1.387$, $b = 1.5$ and $c = 1.063$.

2.4 LQQ

The “linear quadratic quadratic” ψ -function, or *lqq* for short, was proposed by [Koller and Stahel \(2011\)](#). It is defined as,

$$\psi_{b,c,s}(x) = \begin{cases} x & |x| \leq c \\ \text{sign}(x) \left(|x| - \frac{s}{2b} (|x| - c)^2 \right) & c < |x| \leq b + c \\ \text{sign}(x) \left(c + b - \frac{bs}{2} + \frac{s-1}{a} \left(\frac{1}{2} \tilde{x}^2 - a\tilde{x} \right) \right) & b + c < |x| \leq a + b + c \\ 0 & \text{otherwise,} \end{cases}$$

where

$$\tilde{x} := |x| - b - c \quad \text{and} \quad a := (2c + 2b - bs)/(s - 1). \quad (3)$$

The parameter c determines the width of the central identity part. The sharpness of the bend is adjusted by b while the maximal rate of descent is controlled by s ($s = 1 - \min_x \psi'(x) > 1$). From (3), the length a of the final descent to 0 is a function of b , c and s .

```
> cT <- rbind(cc1 = .psi.lqq.findc(ms= -0.5, b.c = 1.5, eff=0.95, bp=NA ),
+             cc2 = .psi.lqq.findc(ms= -0.5, b.c = 1.5, eff=NA , bp=0.50))
> colnames(cT) <- c("b", "c", "s"); cT
```

```
      b      c      s
cc1 1.4734061 0.9822707 1.5
cc2 0.4015457 0.2676971 1.5
```

If the minimal slope is set to $-\frac{1}{2}$, i.e., $s = 1.5$, and $b/c = 3/2 = 1.5$, the constants for 95% efficiency of the regression estimator are $b = 1.473$, $c = 0.982$ and $s = 1.5$, and those for a breakdown point of 0.5 of the S-estimator are $b = 0.402$, $c = 0.268$ and $s = 1.5$.

```
> p.psiFun(x., "LQQ", par = c(-.5, 1.5, .95, NA))
```

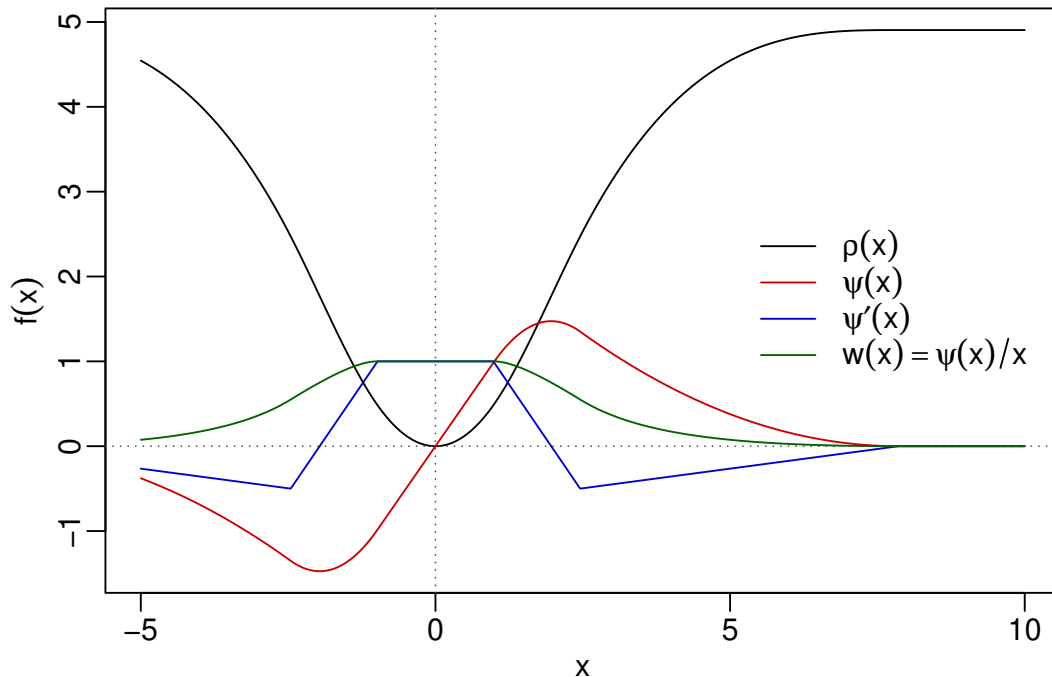


Figure 5: LQQ family of functions using tuning parameters $b = 1.473$, $c = 0.982$ and $s = 1.5$.

2.5 Optimal

The optimal ψ function as given by [Maronna et al. \(2006, Section 5.9.1\)](#),

$$\psi_c(x) = \text{sign}(x) \left(-\frac{\varphi'(|x|) + c}{\varphi(|x|)} \right)_+,$$

where φ is the standard normal density, c is a constant and $t_+ := \max(t, 0)$ denotes the positive part of t .

Note that the `robustbase` implementation uses rational approximations originating from the `robust` package's implementation. That approximation also avoids an anomaly for small x and has a very different meaning of c .

The constant for 95% efficiency of the regression estimator is 1.060 and the constant for a breakdown point of 0.5 of the S-estimator is 0.405.

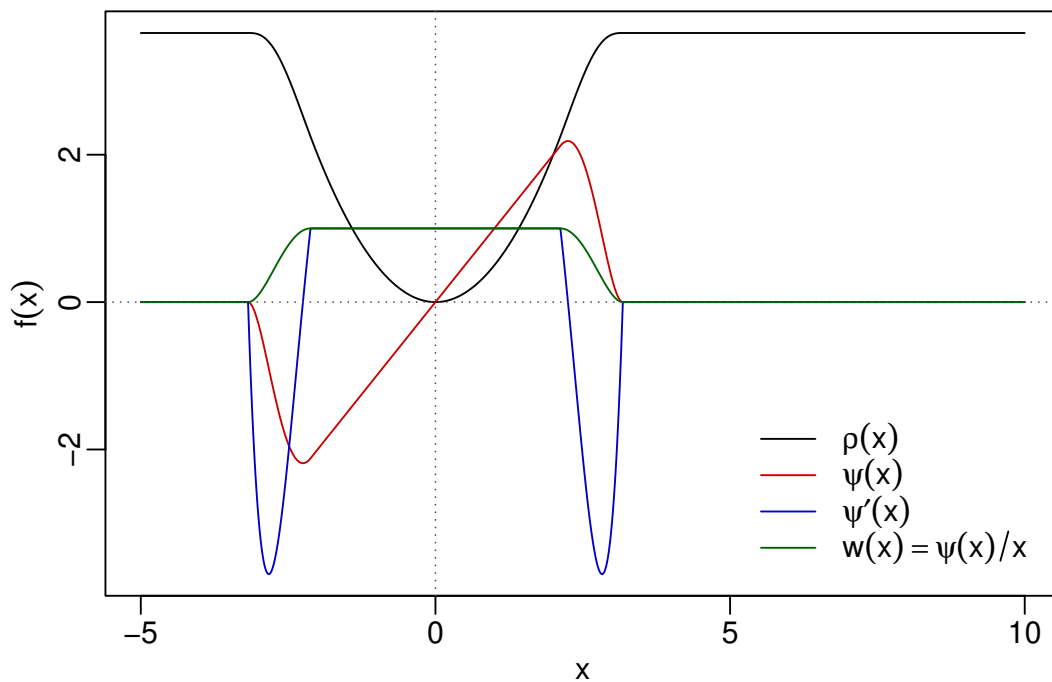


Figure 6: ‘Optimal’ family of functions using tuning parameter $c = 1.06$.

2.6 Welsh

The Welsh ψ function is defined as,

$$\begin{aligned}\tilde{\rho}_k(x) &= 1 - \exp(-(x/k)^2/2) \\ \psi_k(x) &= k^2 \tilde{\rho}'_k(x) = x \exp(-(x/k)^2/2) \\ \psi'_k(x) &= (1 - (x/k)^2) \exp(-(x/k)^2/2)\end{aligned}$$

The constant k for 95% efficiency of the regression estimator is 2.11 and the constant for a breakdown point of 0.5 of the S-estimator is 0.577.

Note that GGW (subsection 2.3) is a 3-parameter generalization of Welsh, matching for $b = 2$, $c = 0$, and $a = k^2$ (see R code there):

```
> ccc <- c(0, a = 2.11^2, b = 2, c = 0, 1)
> (ccc[5] <- integrate(.Mpsi, 0, Inf, ccc=ccc, ipsi = 5)$value) # = rho(Inf)
[1] 4.4521

> stopifnot(all.equal(Mpsi(x., ccc, "GGW"), ## psi[ GGW ](x; a=k^2, b=2, c=0) ==
+ Mpsi(x., 2.11, "Welsh")))## psi[Welsh](x; k)
```

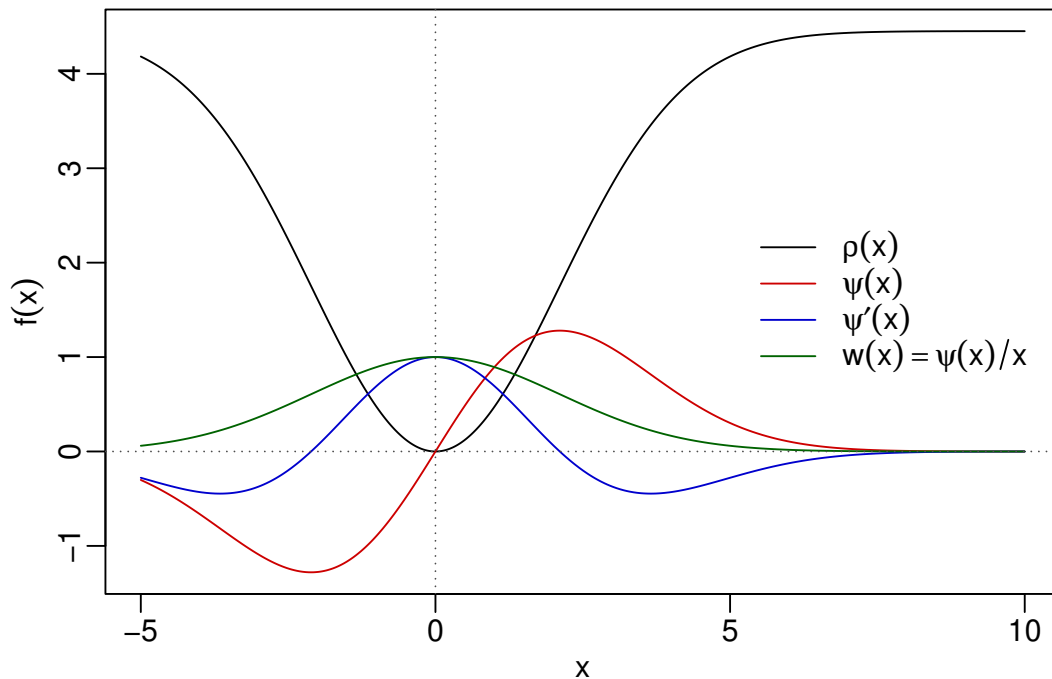


Figure 7: Welsh family of functions using tuning parameter $k = 2.11$.

References

- Hampel, F., E. Ronchetti, P. Rousseeuw, and W. Stahel (1986). *Robust Statistics: The Approach Based on Influence Functions*. N.Y.: Wiley.
- Huber, P. J. (1964). Robust estimation of a location parameter. *Ann. Math. Statist.* 35, 73–101.
- Koller, M. and W. A. Stahel (2011). Sharpening wald-type inference in robust regression for small samples. *Computational Statistics & Data Analysis* 55(8), 2504–2515.
- Maronna, R. A., R. D. Martin, and V. J. Yohai (2006). *Robust Statistics, Theory and Methods*. Wiley Series in Probability and Statistics. John Wiley & Sons, Ltd.